Nonparametric Bayesian volatility learning under microstructure noise

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Team







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Summary

Problem formulation

Problem

• Model: Stochastic differential equation (SDE)

$$dX_t = b(t, X_t)dt + s(t)dW_t, \quad X_0 = x_0, \quad t \in [0, T].$$

 Task: Learn the volatility function s from noisy observations on X at times 0 < t₁ < ... < t_n = T:

$$Y_i = X_{t_i} + V_i.$$

The V's are unobserved stochastic disturbances.

• Important problem in *finance*; data subject to *microstructure noise* and not on a *uniform* time grid.

Sources of inspiration



- *Minimalistic* in its assumptions on the volatility function, which in particular can be a stochastic process.
- Intuitive to understand. Ingredients are well-known techniques.
- Posterior inference via Gibbs sampler.

Learning procedure

- *Closed form expression* for the posterior distribution not available; numerically intractable likelihood.
- Remedy: misspecification on purpose
- Alternative: Data augmentation device and some carefully chosen MCMC sampler.

- Model misspecification not necessarily bad, if you know what you are doing; cf. GUGUSHVILI ET AL. '20 and MARTIN ET AL. '18.
- *Misspecify the model purposely* and act as if the drift *b* = 0. Ultimate justification: Girsanov's theorem.
- Then

$$X_{t_i} = X_{t_{i-1}} + U_i.$$

Here

$$U_i = \int_{t_{i-1}}^{t_i} s(t) \mathrm{d}W_t \sim N(0, w_i),$$

and $w_i = \int_{t_{i-1}}^{t_i} s^2(t) dt$. Note that $\{U_i\}$ are independent.

• Simplify the notation: $x_i = X_{t_i}$, $y_i = Y_i$, $u_i = U_i$, $v_i = V_i$. Then

$$x_i = x_{i-1} + u_i,$$

$$y_i = x_i + v_i.$$

This is a linear state-space model.

- Assume that $\{v_i\} \stackrel{iid}{\sim} N(0, \eta_v)$ are independent of the Wiener process W, so that $\{v_i\}$ are independent of $\{u_i\}$.
- The model becomes Gaussian. This is very convenient.
- Notation:

$$\mathcal{X}_n = \{x_i\}, \quad \mathcal{Y}_n = \{y_i\}.$$

 $\bullet~$ Define N~ bins

$$B_k = [t_{m(k-1)}, t_{mk}), \quad k = 1, \dots, N.$$

where n = mN.

• Assume piecewise constant (squared) volatility,

$$s = \sum_{k=1}^{N} \sqrt{\theta_k} \mathbf{1}_{B_k}, \qquad s^2 = \sum_{k=1}^{N} \theta_k \mathbf{1}_{B_k}.$$

• Assume $\{\theta_k\}$ form an inverse Gamma Markov chain...

• Fix hyperparameters α_1 , α and define a Markov chain on $\theta_1, \zeta_2, \theta_2, \dots, \zeta_k, \theta_k, \dots, \zeta_N, \theta_N$ with (latent) auxiliary variables ζ_k , $k = 2, \dots, N$, and transition densities specified via

$$\theta_1 \sim \operatorname{IG}(\alpha_1, \alpha_1),$$

$$\zeta_{k+1} \mid \theta_k \sim \operatorname{IG}(\alpha, \alpha \theta_k^{-1}),$$

$$\theta_{k+1} \mid \zeta_{k+1} \sim \operatorname{IG}(\alpha, \alpha \zeta_{k+1}^{-1}).$$

• *Mild smoothing/regularisation* through the prior-induced positive correlation of the volatility function across close bins.

Realisations of volatility under the prior



Figure 1: Realisations of $\{\theta_k\}$ with $\alpha = 30$ for N = 400. If $\log \theta_k$ looks like cW_t to you... you are right.

Realisations of volatility under the prior



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- Drift $b \equiv 0$.
- IGMC for $\{\theta_i\}$.
- Diffuse prior for α ; improper for α_1 .
- Noise variance $\eta_v \sim IG$.

• Sequential updates:

$$\begin{split} \zeta_2, \dots, \zeta_N \mid \mathcal{X}_n, \ \alpha, \theta_1, \dots, \theta_N \quad (\text{conjugate}) \\ \theta_1, \dots, \theta_N \mid \mathcal{X}_n, \alpha, \zeta_2, \dots, \zeta_N \quad (\text{conjugate}) \\ \alpha \mid \theta_1, \zeta_2, \dots, \zeta_N, \theta_N \quad (\text{generic MH}) \\ \mathcal{X}_n \mid \mathcal{Y}_n, \eta_v, \theta_1, \dots, \theta_N \quad (\text{Kalman-based}) \\ \eta_v \mid \mathcal{V}_n \quad (\text{conjugate}) \end{split}$$

Numerical experiment

- Widely used stochastic volatility model (HESTON '93).
- Price process \boldsymbol{S} evolving according to the SDE

$$\mathrm{d}S_t = \mu S_t \mathrm{d}t + \sqrt{Z_t} S_t \mathrm{d}W_t,$$

where the process Z follows the CIR (COX ET AL. '85) or square root process,

$$\mathrm{d}Z_t = \kappa(\theta - Z_t)\mathrm{d}t + \sigma\sqrt{Z_t}\mathrm{d}B_t.$$

W and B are correlated Wiener processes with correlation ρ .

• By Itō's formula, logarithm $X_t = \log S_t$ obeys the diffusion equation with volatility

$$s(t) = \sqrt{Z_t}.$$

• In the Heston model the volatility function *s* is random, with its law not known in closed form.

- High frequency observations on the log-price process X with additive noise V_i ~ N(0, η). No further knowledge of the data generation mechanism.
- Simulated path using "realistic" values (cf. Table 1 in HESTON '93)

$$\mu = 0.05, \quad \kappa = 7, \quad \theta = 0.04, \quad \sigma = 0.6, \quad \rho = -0.6.$$

• Noise variance $\eta_v = 10^{-6}$ yielding a meaningful signal-to-noise ratio.

Heston model iv



Figure 2: Posterior mean and pointwise 95% credible band for N = 80 bins. True volatility function plotted in red; the black step function is the posterior mean.

Application to real data

- Volatility learning from the high frequency exchange rate data.
- EUR/USD tick data (bid prices) for 2 March 2015, subsampled by retaining every 10th observation. A total of $n = 13\,025$ observations (subsampling accounts for the assumption of independent additive measurement noise).
- Log-transform the observed time series and assume the additive measurement error model.

Exchange rate data ii



Figure 3: Top: Logarithm of the EUR/USD exchange rate data for 2 March 2015. Bottom: Posterior mean (black curve) and pointwise 95% credible band (blue band) for the volatility function.

Summary

- A conceptually straightforward, clever nonparametric Bayesian method for volatility learning under microstructure noise.
- Available as julia package; see SCHAUER AND GUGUSHVILI '20.
- Partial fulfillment of prediction in GODSILL ET AL. '07 that some ideas developed originally in the context of audio and music processing "will also find use in other areas of science and engineering, such as financial or biomedical data analysis".

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