

Prediction of future mortality

Mathias Lindholm

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The presented material is based primarily on joint work with
Patrik Andersson, Uppsala University

in particular the paper

“Mortality forecasting using a Lexis-based state-space model”

To appear in *Annals of Actuarial Science*

Outline of the presentation

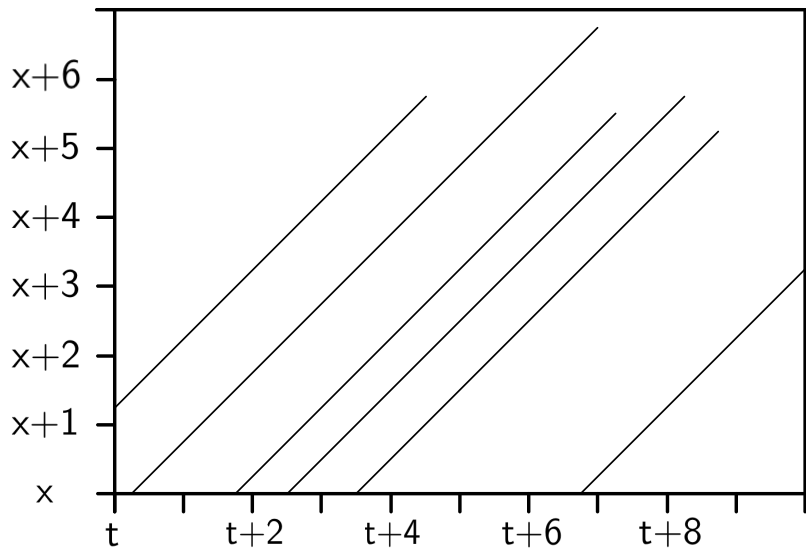
- ▶ What do we observe?
- ▶ Discuss how to go from observed mortality to predicting future mortality
- ▶ From Lee-Carter to state-space models
- ▶ Some words about ongoing work and extensions

Some notation:

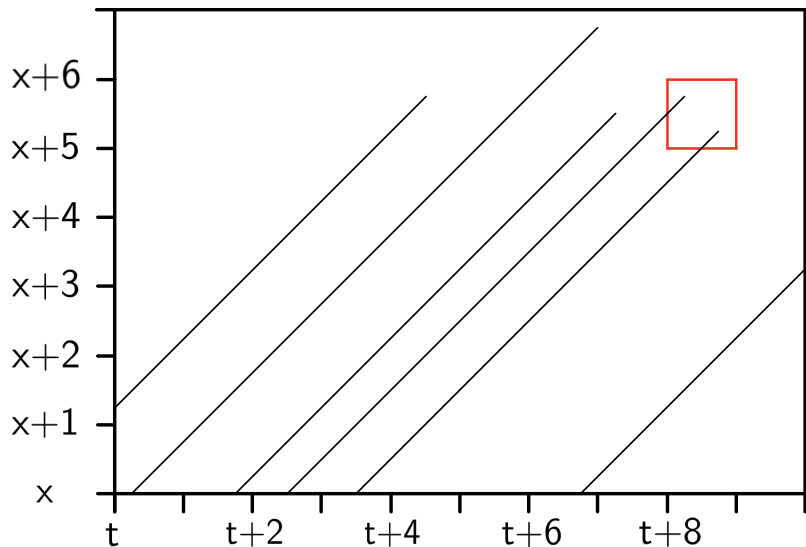
- ▶ all vectors are column vectors
- ▶ we will usually not define the dimension of vectors and matrices – should be clear from context
- ▶ X corresponds to a random quantity, x corresponds to an observation
- ▶ $x_{0:n} = (x_0, \dots, x_n)$
- ▶ all Greek letters corresponds to model parameters

What we observe and structuring of data – Lexis diagrams

Lexis diagrams



Lexis diagrams



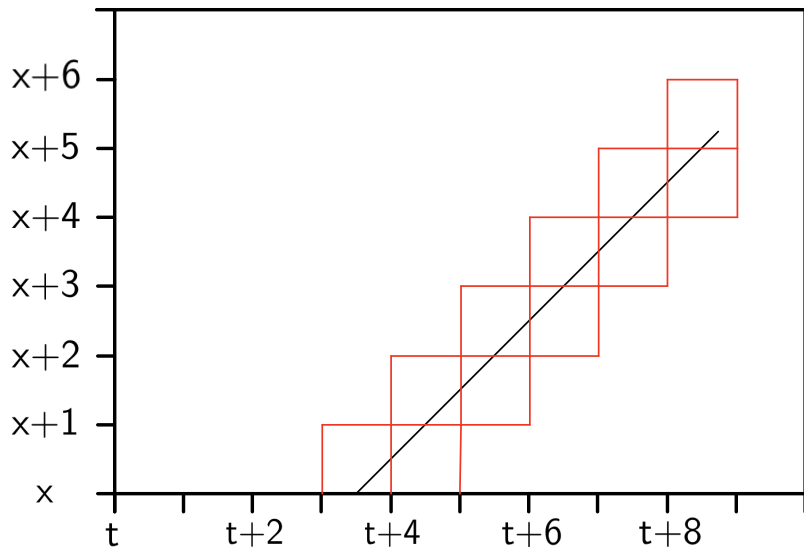
Lexis diagrams: models and likelihood

Assumptions:

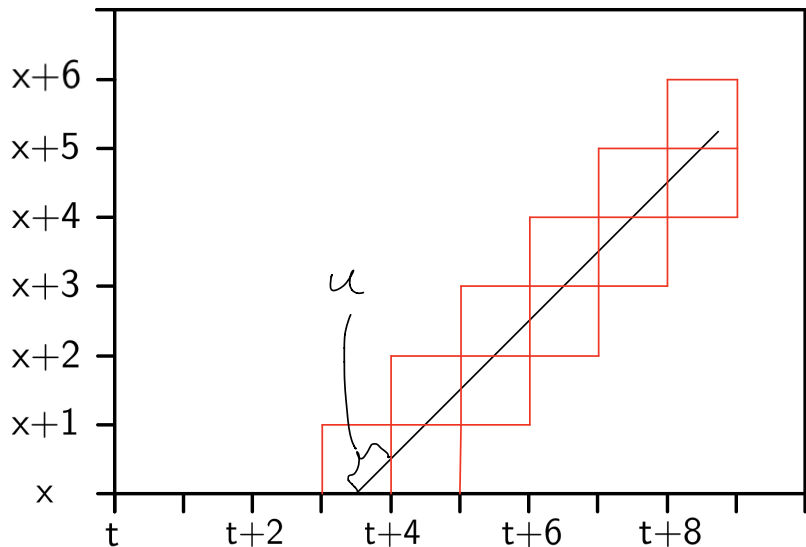
- ▶ all individuals are independent
- ▶ the mortality rate is constant within yearly Lexis squares

Start with a single individual!

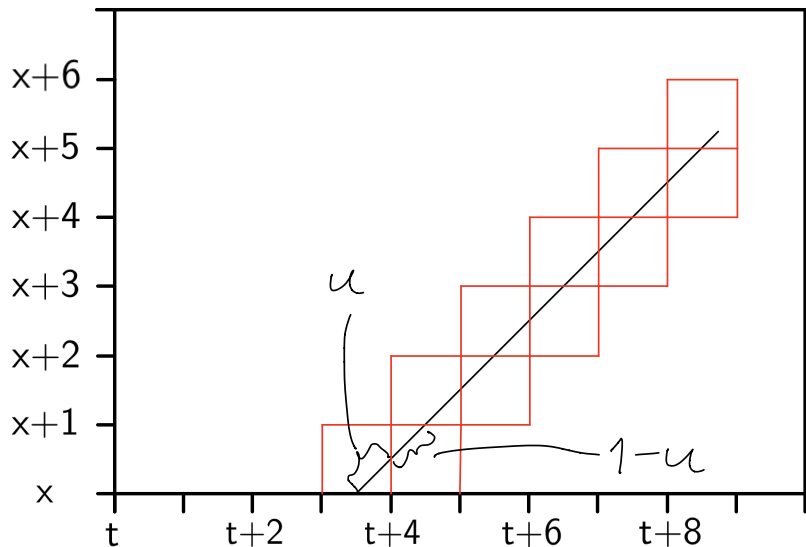
Lexis diagrams: models and likelihood – single individual



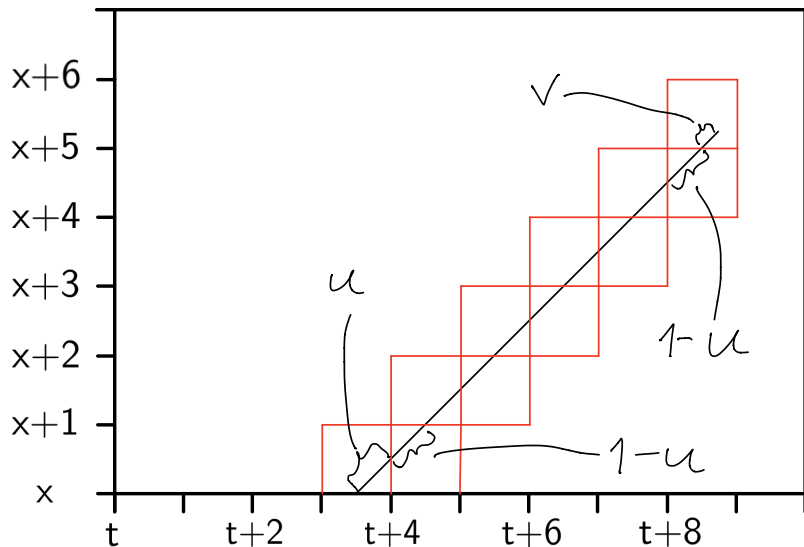
Lexis diagrams: models and likelihood – single individual



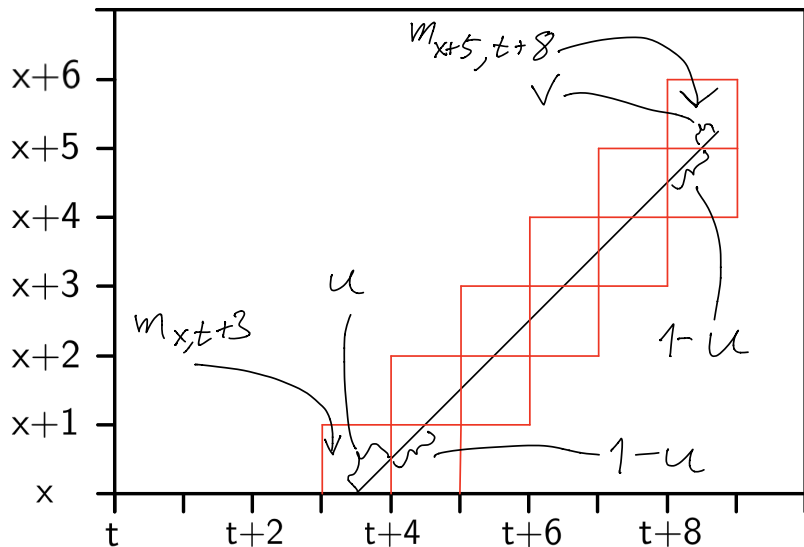
Lexis diagrams: models and likelihood – single individual



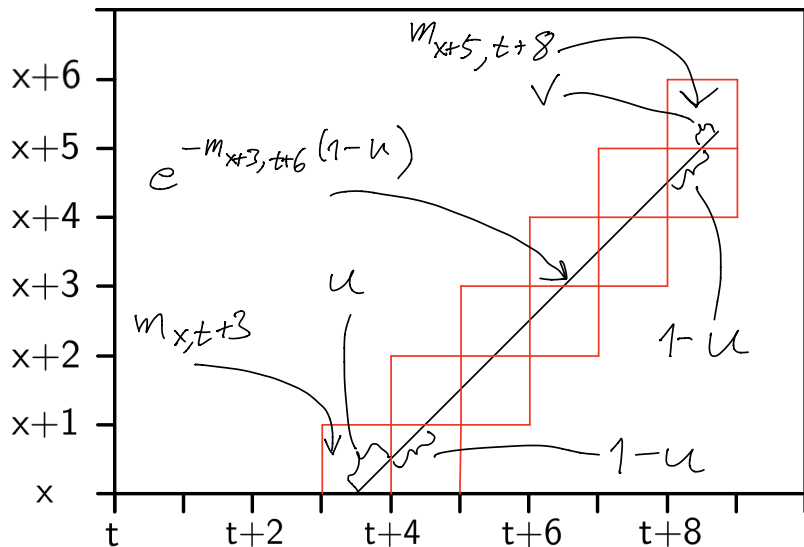
Lexis diagrams: models and likelihood – single individual



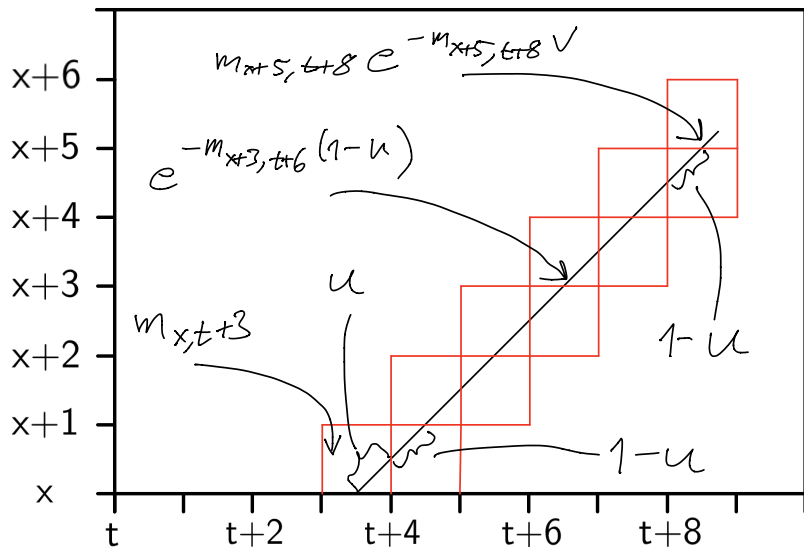
Lexis diagrams: models and likelihood – single individual



Lexis diagrams: models and likelihood – single individual



Lexis diagrams: models and likelihood – single individual

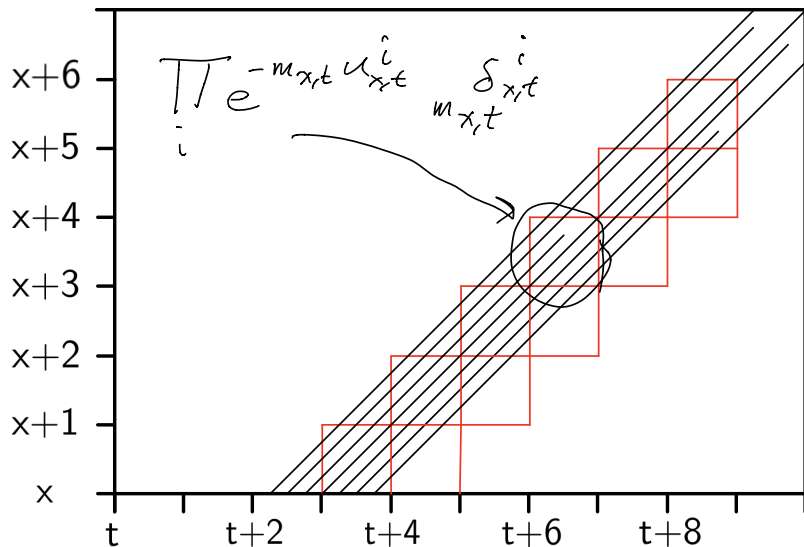


Lexis diagrams: models and likelihood – single individual

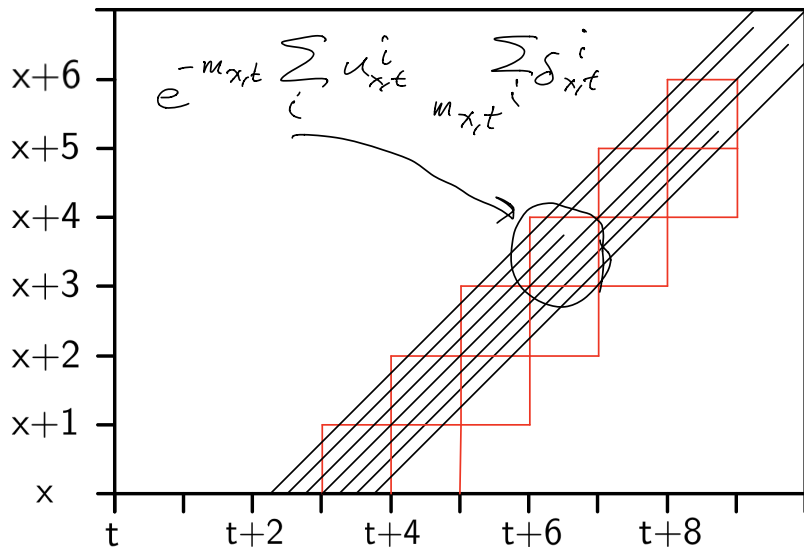
Single individual: multiply contribution from all Lexis squares visited by the individual

Single square, many individuals?

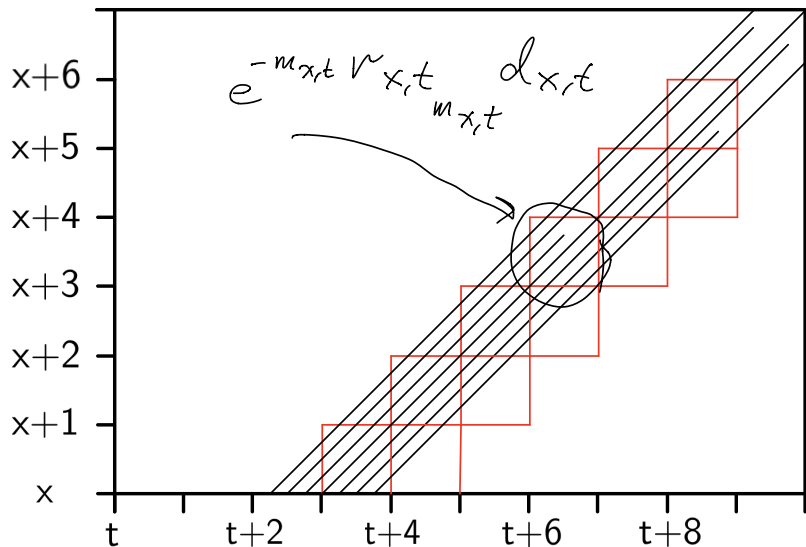
Lexis diagrams: models and likelihood – single square



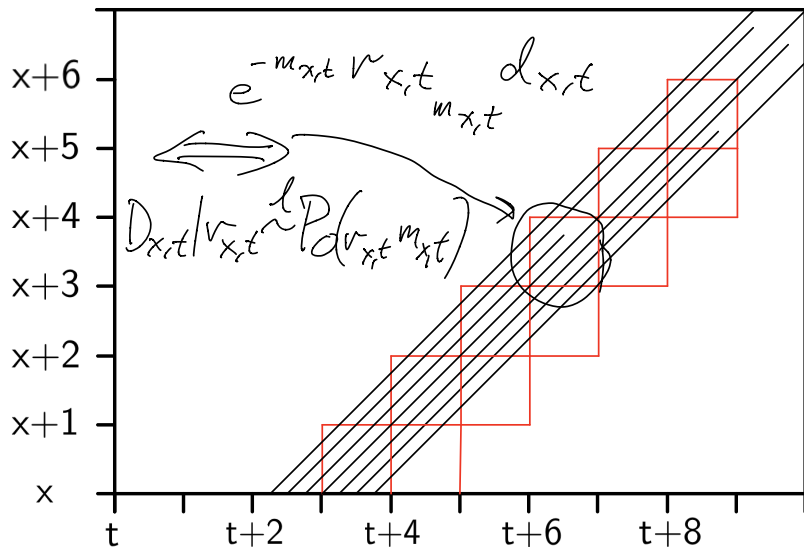
Lexis diagrams: models and likelihood – single square



Lexis diagrams: models and likelihood – single square



Lexis diagrams: models and likelihood – single square



Lexis diagrams: models and likelihood

Lemma 1

The log-likelihood for the total population, assuming independence between individuals and piecewise constant $m_{x,t}$, is,

$$l(\mathcal{M}) = \sum_{(x,t) \in \mathcal{S}} (d_{x,t} \log m_{x,t} - r_{x,t} m_{x,t}), \quad (1)$$

where \mathcal{S} denotes the set of all observed Lexis squares, $\mathcal{M} = \{m_{x,t} \mid (x,t) \in \mathcal{S}\}$, $d_{x,t}$ corresponds to the observed number of deaths in (x,t) and $r_{x,t}$ corresponds to the observed exposure-to-risk in (x,t) .

Consequence of Lemma 1:

$$\hat{m}_{x,t} = \frac{d_{x,t}}{r_{x,t}},$$

$\hat{m}_{x,t}$ is often referred to as the “force of mortality” or “mortality rate”.

Lexis diagrams: models and likelihood – prediction

...but we want to predict future mortality rates...

...and note that Lemma 1 **provides no structure between** $m_{x,t}$ s

We need additional structure!

Can be done in a lot of different ways!

Lexis diagrams: models and likelihood

The “Lee-Carter approach”, see e.g. [3]:

1. Estimate $\hat{m}_{x,t}$ for all $(x, t) \in \mathcal{S}$
2. Given $\{\hat{m}_{x,t}\}$, fit a stochastic process to the estimated $\hat{m}_{x,t}$ s

Sometimes a “post fitting adjustment” is done, due to problems caused by the two-stage procedure

Lexis diagrams: models and likelihood

In somewhat more detail:

1. Estimate $\hat{m}_{x,t}$ for all $(x, t) \in \mathcal{S}$ and let Y be the $m \times n$ matrix defined as

$$Y_{x,t} = \log \hat{m}_{x,t}$$

2. Let \bar{y} denote the $m \times 1$ vector with age-wise estimated average mortality rates, and make the following SVD-approximation:

$$(Y - \bar{y})_{x,t} \approx \beta_x \kappa_t$$

3. Fit a (Gaussian) stochastic process to the estimated $\hat{\kappa}_t$ s

Note: 2. corresponds to **dimension reduction**

Lexis diagrams: models and likelihood

Note the close resemblance between the above procedures and the following state space model (see e.g. [1, 2]):

$$\begin{cases} Y_t = \mu + \beta K_t + U_t, & U_t \sim N(0, \sigma_u^2 I) \\ K_t = \alpha + K_{t-1} + V_t, & V_t \sim N(0, \sigma_v^2) \end{cases}$$

where $\mu, \beta, \alpha, \sigma_u$, and σ_v are unknown parameters and the K_t s correspond to **unobservable states**

Still, note that Y is a matrix based on **estimated** mortality rates!

Lexis diagrams: models and likelihood

Note that

- ▶ the Lee-Carter model is a model for **mortality rates**
- ▶ we **observe death counts**

Taking the above into account (Andersson & ML, [4]):

Use a Poisson state space model for **death counts**

Poisson state space models

The Poisson state space model

$$\left\{ \begin{array}{l} D_{i,t} \mid M_{i,t}, r_{i,t} \sim \text{Po}(r_{i,t} M_{i,t}) \\ M_{i,t} = \exp \{ (\Upsilon X_t)_i \} \\ X_{t+1} = \Gamma X_t + \mu + V_t, \quad V_t \sim \text{N}(0, \Sigma) \\ X_0 \sim \text{N}(\mu_0, \Sigma_0) \end{array} \right. \quad (*)$$

Note that

- ▶ Υ corresponds to the dimension reduction from m to p (a.k.a. “EPCA” or “GPCA”), where $X_t \in \mathbb{R}^p$, where m is the number of age groups
- ▶ $M_{i,t}$ is defined in terms of the standard link-function for a Poisson GLM and that $\log M_{i,t} \approx Y_{i,t}$

Poisson state space models

Note that, given estimated parameters and state vectors, $x_{0:t}$, it is easy to

- ▶ decompose the in-sample variation of

$$\widehat{M}_{i,t} := \frac{D_{i,t}}{r_{i,t}}$$

in terms of variation from $M_{i,t} = \exp\{(\Upsilon X_t)_i\}$ and variation from $D_{i,t}$ conditional on $M_{i,t}$ and $r_{i,t}$

- ▶ project future mortality rates (and death counts) by first sampling a trajectory $x_{0:n}$

Estimation

Following [4], estimation of **parameters** and **state vectors** is done using Stochastic Approximation EM + particle filter techniques

- ▶ in a first step Υ is estimated using a Poisson likelihood
- ▶ given $\hat{\Upsilon}$, the remaining model parameters, Γ , μ , and Σ , and state-vectors are estimated
- ▶ the SAEM algorithm makes use of that we have explicit sufficient statistics for the remaining model parameters these sufficient statistics are updated iteratively using a weighted averaging
- ▶ in each SAEM update state-vectors are sampled using the Forward Filtering Backward Smoothing algorithm given the latest parameter updates

Model evaluation and forecasting

We will split our data into two parts

- ▶ in-sample “training data”
- ▶ out-of-sample “validation data”

Recall that model (★) allows us to

- ▶ predict future **mortality rates**
- ▶ (based on simulations) calculate

$$M_{i,t} = \frac{D_{i,t}}{r_{i,t}} \quad (\text{force of mortality}) \quad (2)$$

By using (2) for model validation we will capture both

- (i) the population (“Poisson”) variation
- (ii) the variation from the mortality rates

Model evaluation and forecasting

Concerning model selection, Andersson & ML, [4], suggest an R^2 -type measure that

- ▶ is based on using the likelihood
- ▶ works both in-sample and out-of-sample

Numerical illustrations

Numerical illustrations

We will use the following special case of model (\star):

$$\left\{ \begin{array}{l} D_{i,t} \mid M_{i,t}, r_{i,t} \sim \text{Po}(r_{i,t} M_{i,t}) \\ M_{i,t} = \exp \{ (\Upsilon X_t)_i \} \\ X_{t+1} = \Gamma_{xx} X_t + \Gamma_{xy} Y_t + \mu_x + V_t, \quad V_t \sim \text{N}(0, \Sigma_x) \\ Y_{t+1} = \Gamma_{yy} Y_t + \mu_y + U_t, \quad U_t \sim \text{N}(0, \Sigma_y) \\ X_0 \sim \text{N}(\mu_0, \Sigma_0) \\ Y_0 \sim \text{N}(\mu_0, \Sigma_0) \end{array} \right.$$

Numerical illustrations

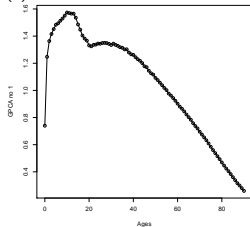
We will

- ▶ use Swedish male data from 1930–2016 with ages 0–90
- ▶ use 1–3 GPCA components
- ▶ illustrate the role of the Poisson part of the state space modelling

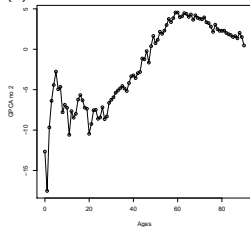
To start off: SWE male population, estimation 1930–1960, validation 1961–2016

(More examples, and more on model selection, can be found in Andersson & ML, [4])

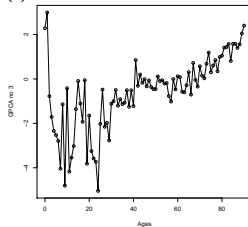
(a) 1st GPCA



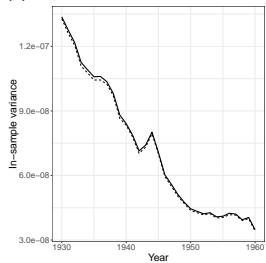
(b) 2nd GPCA



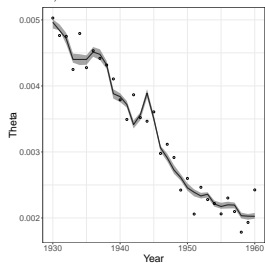
(c) 3rd GPCA



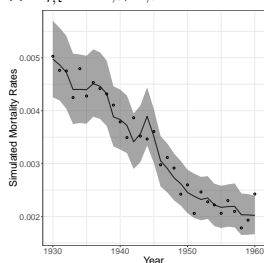
(d) Population var vs total var



(e) $M_{i,t} = \exp(\Upsilon X_t)_i$

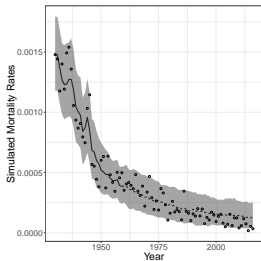


(f) $M_{i,t}^{\text{sim}} = D_{i,t}/r_{i,t}$

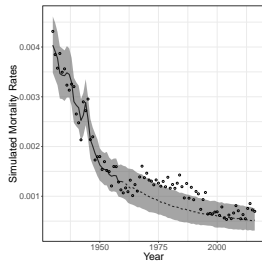


SWE males, 3 GPCA

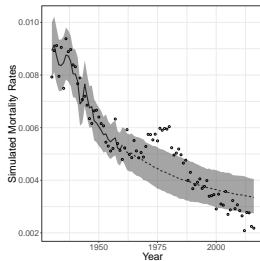
(a) Fit 1930–1960, 10 yrs



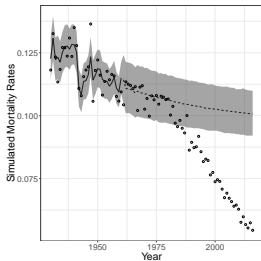
(b) Fit 1930–1960, 30 yrs



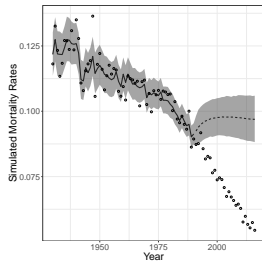
(c) Fit 1930–1960, 50 yrs



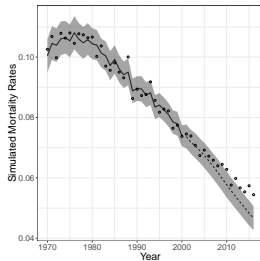
(d) Fit 1930–1960, 80 yrs



(e) Fit 1930–1990, 80 yrs



(f) Fit 1970–2000, 80 yrs



Concluding remarks and extensions

We have





- ▶ introduced a flexible Poisson state space model which allows for dimension reduction...
...that avoids two-step estimation procedures
- ▶ illustrated the importance of explicitly capturing the Poisson part of the variation

...but

- ▶ we have used SAEM + particle filter techniques
– other choices?
- ▶ one could consider allowing for even more flexibility
– neural networks? ...estimation!?

...this is work in progress

References I

-  Piet De Jong and Leonie Tickle.
Extending lee–carter mortality forecasting.
Mathematical Population Studies, 13(1):1–18, 2006.
-  James Durbin and Siem Jan Koopman.
Time series analysis by state space methods.
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-  Ronald D Lee and Lawrence R Carter.
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-  Andersson Patrik and Mathias Lindholm.
Mortality forecasting using a lexis-based state-space model.
To appear in Annals of Actuarial Science, 2020.

Appendix

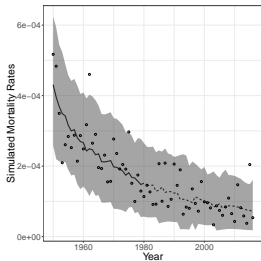
Female populations in SWE and US

Essentially similar results, will focus on ages 10, 40, and 80

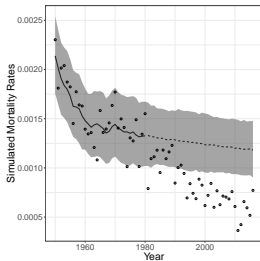
- ▶ SWE females, fit 1950–1980, 2 GPCA
- ▶ US females, fit 1950–1980, 3 GPCA

Fit 1950–1980

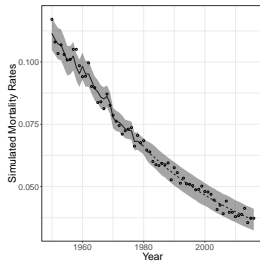
(a) SWE females, 2 GPCA, 10 yrs



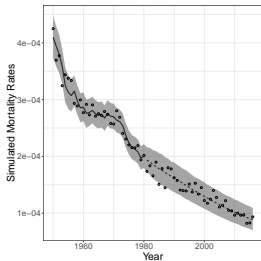
(b) 40 yrs



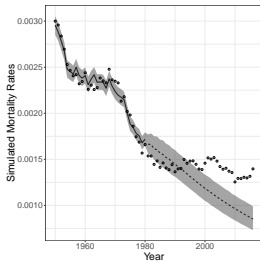
(c) 80 yrs



(d) US females, 3 GPCA, 10 yrs



(e) 40 yrs



(f) 80 yrs

