

# Using modern statistical methodology for validating and reporting Patient Reported Outcomes

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## ① Indirect measurement

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- 2 Ordinal regression models with random effects
  - Item Response Theory (IRT) models
  - Probit models

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- 4 Summary

- 1 Binet-Simon scale, first intelligence test
  - Designed for children, measured a child's 'mental age' which compared the child with his or her age group (Binet, L'Année Psychologique, vol. 12, 191-244, 1905)



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- latent variable, can be changed

## ① Karnofsky performance status

- 0 ('Dead'), 10 ('Moribund'), 20 ('Very sick'), ..., 90 ('Able to carry on normal activity; minor symptoms'), 100

(Karnofsky, Burchenal JH, In: MacLeod CM (ed.): Evaluation of chemotherapeutic agents. New York: Columbia University Press, pp 191-205, 1949)

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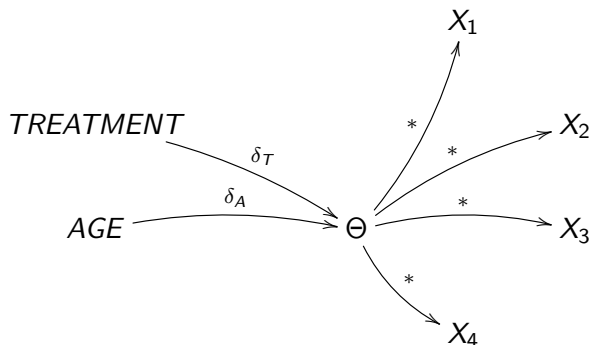
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## 1 PRO's

- Generic (e.g. SF36) or disease specific (e.g. FACT)
- Sub scales consisting of ordinal items
- Often reported as standardized (e.g. zero to 100) mean scores
- Often analyzed using linear models or nonparametric statistics

# Indirect measurement



Latent variable  $\Theta$ ,  $X_1, \dots, X_4$  items. [\*: monotone association]

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Item parameters  $\bar{\beta}_i = (\beta_{i0}, \beta_{i1}, \beta_{i2}, \dots)$

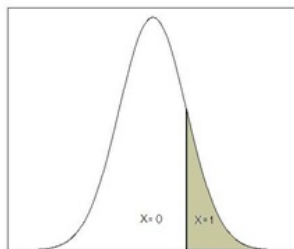
$$P(X_i = x | \Theta = \theta) = \frac{\exp(x\theta + \beta_{ix})}{K_i(\theta)} \quad (x = 0, 1, \dots, m_i), \quad (1)$$

where  $\beta_{i0} = 0$  for convenience and

$$K_i(\theta) = \sum_{h=0}^{m_i} \exp(h\theta + \beta_{ih})$$

# Alternative models for item response $X_i$

- Change linear predictor: replace  $x\theta + \beta_{ix}$  with  $\alpha_j(x\theta + \beta_{ix})$
- Probit model

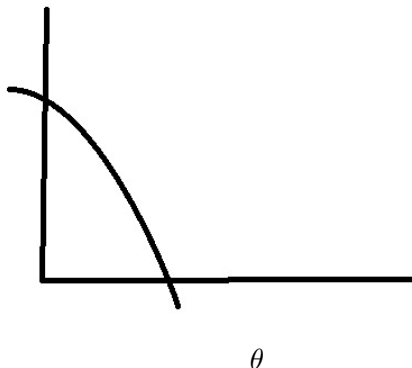


- Regardless of these choices
  - Conditional probability of each response option given  $\theta$



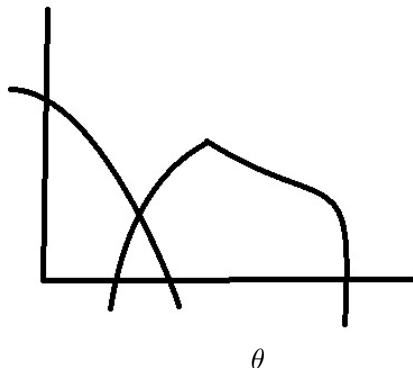
# Conditional probability of each response option given $\theta$

$$P(X = 0 | \Theta = \theta)$$



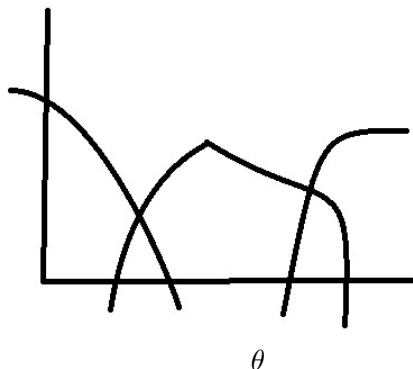
# Conditional probability of each response option given $\theta$

$$P(X = 1 | \Theta = \theta)$$



# Conditional probability of each response option given $\theta$

$$P(X = 2 | \Theta = \theta)$$



# Model for response vector $\bar{X} = (X_i)_{i \in I}$

Probability of observing  $\bar{x} = (x_i)_{i \in I}$  is

$$P(\bar{X} = \bar{x} | \Theta = \theta) \stackrel{*}{=} \prod_{i \in I} P(X_i = x_i | \Theta = \theta) \quad (2)$$

$$= \frac{\exp(\theta \sum_{i \in I} X_i + \sum_{i \in I} \beta_{ix_i})}{\prod_{i \in I} K_i(\theta)} \quad (3)$$

Marginal probability

$$P(\bar{X} = \bar{x}) = \int P(\bar{X} = \bar{x} | \Theta = \theta) \phi(\theta) d\theta \quad (4)$$

$$= \int \frac{\exp(\theta \sum_{i \in I} X_i + \sum_{i \in I} \beta_{ix_i})}{\prod_{i \in I} K_i(\theta)} \varphi(\theta) d\theta \quad (5)$$

[\*: assumption: 'local' independence]

Subjects  $v = 1, \dots, N$ ,  $T_v$  treatment indicator,  $A_v$  age

$$\theta_v = \delta_0 + \delta_T T_v [+ \delta_A A_v] + \epsilon_v \quad (6)$$

$$L(\bar{\beta}, \delta_0, \delta_T, \delta_A) = \prod_{v=1}^N \int P(\bar{X} = \bar{x} | \Theta = \delta_0 + \delta_T T_v + \epsilon_v) \varphi(\epsilon) d\theta \quad (7)$$

Implementation: PROC NLMIXED

Set of items split into two disjoint sets

$$I = I_1 \cup I_2, \quad (8)$$

with items in  $I_1$  and  $I_2$  measuring latent variables  $\theta_1$  and  $\theta_2$  respectively.

$$P(\bar{X} = \bar{x} | \bar{\Theta} = \bar{\theta}) = \frac{\exp(r_1\theta_1 + r_2\theta_2 + \sum_{i \in I} \eta_{ix_i})}{K(\bar{\theta})} \quad (9)$$

where  $r_d = \sum_{i \in I_d} x_i$ ,  $d = 1, 2$  and

$$K(\bar{\theta}) = \left( \prod_{i \in I_1} K_i(\theta_1) \right) \left( \prod_{i \in I_2} K_i(\theta_2) \right), \quad \bar{\theta} = \begin{bmatrix} \theta_1 \\ \theta_2 \end{bmatrix}$$

- (i)  $\theta_1$  and  $\theta_2$  are distinct, but correlated latent variables,
- (ii)  $\theta_1$  and  $\theta_2$  represent repeated measurements of the same latent variable

Test invariance:  $(\bar{\beta}_i)_{i \in I_1}$  and  $(\bar{\beta}_i)_{i \in I_2}$  identical.

- $\mu$  average shift in latent variable, latent correlation:

$$\bar{\theta} \sim N \left( \begin{bmatrix} 0 \\ \mu \end{bmatrix}, \Sigma \right) \quad (10)$$

- baseline means constrained to be equal across randomization groups (Liu et al, Statist. Med. vol. 28, 2509-2530, 2009):

$$\bar{\theta}_v = \begin{bmatrix} \delta_0 \\ \delta_0 \end{bmatrix} + \begin{bmatrix} 0 \\ \delta_T T_v \end{bmatrix} + \begin{bmatrix} \epsilon_{v1} \\ \epsilon_{v2} \end{bmatrix}, \quad \begin{bmatrix} \epsilon_{v1} \\ \epsilon_{v2} \end{bmatrix} \sim N \left( \begin{bmatrix} 0 \\ 0 \end{bmatrix}, \Sigma \right) \quad (11)$$

- Implementation: PROC NL MIXED

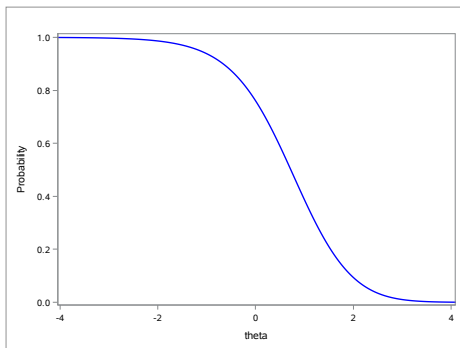


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- 10-item scale, measures limitations
- Items endorsed on 3-point scale ('Yes, Limited A Lot', 'Yes, Limited A Little', 'No, Not Limited At All')
  - Vigorous Activities
  - $\vdots$
  - Walking one block
  - Bathing or dressing yourself

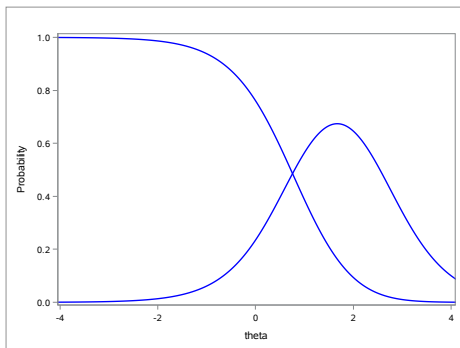
# PF: Conditional probability given $\theta$

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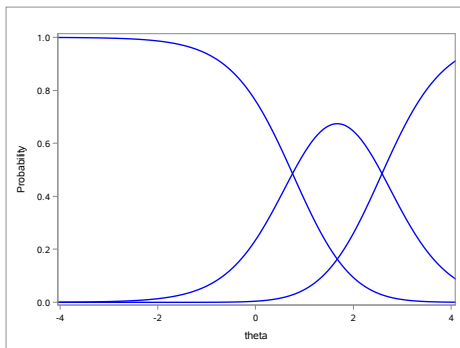
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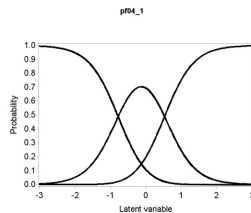
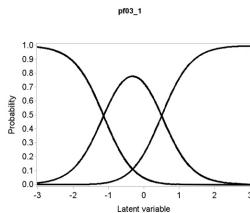
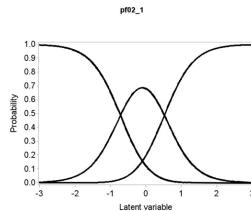
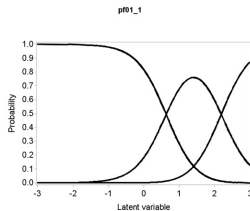


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# Conditional probability given $\theta$



- Widely used measure of vitality
- 4-item scale, (two positively, two negatively-worded items)
  - Did you feel full of pep?
  - Did you have a lot of energy?
  - Did you feel worn out?
  - Did you feel tired?
- Items endorsed on 6-point scale ranging from 'None of the time' to 'All of the time'

# PF and VT sub scales from SF36 - linear models

Data from RCT to determine if patients with acute leukemia can benefit by a structured and supervised counseling and exercise program.

		Baseline			Twelve weeks			Change baseline to twelve weeks				Group comparison			
		N	Mean	SD	N	Mean	SD	N	D	95%	CI	D	95%	CI	p
SF36 Physical Functioning (PF)	A	34	61.9	27.5	32	82.0	21.8	32	19.1	10.0	28.1	10.4	-3.8	24.6	0.15
	B	36	66.7	24.2	30	72.7	26.3	30	8.6	-2.8	20.1				
SF36 Vitality (VT)	A	34	45.4	21.4	32	63.4	19.4	32	17.7	9.8	25.5	11.6	-0.6	23.8	0.06
	B	36	54.6	21.9	30	58.1	25.1	30	6.1	-3.7	15.9				



# PF sub scale from SF36

	T-test			IRT		
	D	(95% CI)	p	$\hat{\delta}_T$	(95% CI)	p
PF	10.4	(-3.8 to 24.6)	0.15	0.56	(-0.11 to 1.22)	0.10

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Power

N=70

N=100

N=100

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Power

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~ 35%

~ 42%

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N=100

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PF	10.4	(-3.8 to 24.6)	0.15	0.56	(-0.11 to 1.22)	0.10

## Power

N=70	~ 35%	~ 42%
N=100	~ 42%	~ 49%
N=130		

# PF sub scales from SF36 - linear models

	T-test			IRT		
	D	(95% CI)	p	$\hat{\delta}_T$	(95% CI)	p
PF	10.4	(-3.8 to 24.6)	0.15	0.56	(-0.11 to 1.22)	0.10

## Power

N=70	~ 35%	~ 42%
N=100	~ 42%	~ 49%
N=130	~ 53%	~ 64%

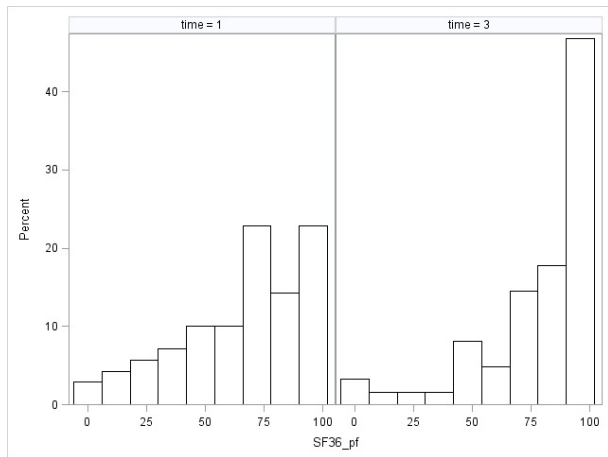
# What about the VT sub scale?

	T-test			IRT		
	D	(95% CI)	p	$\hat{\delta}_T$	(95% CI)	p
PF	10.4	(-3.8 to 24.6)	0.15	0.56	(-0.11 to 1.22)	0.10
VT	11.6	(-0.6 to 23.8)	0.06			

# What about the VT sub scale?

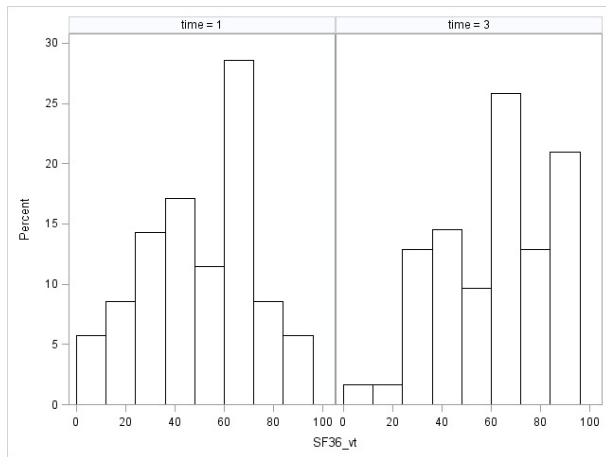
	T-test			IRT		
	D	(95% CI)	p	$\hat{\delta}_T$	(95% CI)	p
PF	10.4	(-3.8 to 24.6)	0.15	0.56	(-0.11 to 1.22)	0.10
VT	11.6	(-0.6 to 23.8)	0.06	0.37	(-0.12 to 0.86)	0.14

# PF sub scale from SF36





# VT sub scale from SF36



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- Nonlinear random effects models yield increased power for skewed data
- Can specify model for the (unobserved) variable of interest
- Can quantify the uncertainty on change scores estimated for individual patients
- Implementation: PROC NL MIXED